

Peer-reviewed research

Which Components of Oil Prices Predict Inflation Rate? New Evidence From a Small Island Country, Fiji

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Keywords: Predictability, Inflation, Oil price, Diesel, Premix

JEL Classifications: C53 Forecasting and Prediction Methods - Simulation Methods, E31 Price Level - Inflation - Deflation, Q43 Energy and the Macroeconomy

https://doi.org/10.46557/001c.129182

Energy RESEARCH LETTERS

Vol. 6, Issue Early View, 2025

This note is about evaluating the role of oil prices and its constituents (namely, diesel, premix, kerosene and motor spirit) in predicting inflation rate. Using time-series monthly data for Fiji unlike for developed markets we do not find any evidence that oil price predicts inflation. We however find strong evidence that motor spirit and premix prices predict inflation by 2.71% and 2.27%, respectively, while diesel predicts inflation by 0.43%.

I. Introduction

Forecasting inflation is the subject of intense research in macroeconomics. The search for successful predictors of inflation is an exercise that has implications for not only inflation targeting but macroeconomic stability. There is a growing literature on forecasting/predicting inflation rates (see Bec & De Gaye, 2016; Gao et al., 2014; Neely, 2015). The motivation for this literature to investigate the oil price inflation nexus is rooted in household expectations of inflation which they revise every time there is an increase in oil prices or its constituents.

With the importance of oil to macroeconomic stability and economic growth, a sub-set of the literature on forecasting inflation treats oil prices as a predictor. Several studies, as a result, show that oil prices predict inflation. See studies such as Salisu et al. (2017) and Salisu and Isah (2018). A feature of these studies is that they are based on either emerging markets or large developed or developing countries. In this literature, therefore, smaller island economies, where inflationary pressures are equally if not more relevant, are almost always ignored. One such country is Fiji Islands, an island economy which is a leader country in the South Pacific, where growing oil prices have exerted fiscal pressures, leading to debt sustainability concerns.

Our contribution to understanding Fiji's macroeconomic stability is from the point of view of inflation. We attempt to understand what precise role oil price and its constituents, such as diesel, premix, kerosene, and motor spirit play in shaping Fiji's inflation rate. This is important to investigate because inflation in Fiji has exceeded the 3-5% inflation target rate (see Figure 1). As noted in Narayan et al.

(2023) the Reserve Bank of Fiji aims to maintain inflation around the 3-5 percent band.

Using a time-series predictability model developed by Westerlund and Narayan (2012, 2015) that deals with data persistency and model endogeneity and heteroskedasticity we document that: (a) oil prices and kerosene do not predict Fiji's inflation rate; diesel prices predict a net 0.43% increase in Fiji's inflation from every one dollar increase in diesel prices and this effect translates over the first 3 months of the price increase; (c) a one dollar increase in the price of premix predicts an increase in Fiji's inflation rate by 2.27%; and (d) a one dollar increase in the price of motor spirits predicts an increase in inflation of 2.71%.

II. Econometric Framework

The time-series predictive regression model has the following form

$$INF_t = \propto +\beta Price_{t-1} + \rho \Delta Price_t + e_{INF,t}$$
(1)

In this relationship, INF is the percentage change in the consumer price index obtained from the Reserve Bank of Fiji and Price are proxies for the oil market namely the crude oil price, diesoline, kerosene, premix and motor spirit. Westerlund and Narayan (2012, 2015) show that the general form of Equation (1) is sufficient to address: (a) any predictor persistence, which we estimate using a first order autoregressive model obtained using ordinary least squares, (b) endogeneity which we obtain the regressing the residuals from a regression of *INF* on one period lagged *Price* variable on the residuals from a 6th order autoregressive model of the *Price* variable and the null hypothesis of no endogeneity is associated with the resulting slope co-

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Figure 1. Annualized inflation for Fiji, 2011 to 2022

Figure 1 represents the annualized inflation rate for Fiji computed from the consumer price index over the period 2011 to 2022.

efficient from the residual-regressed model., and (c) heteroskedasticity, which we address by estimating the predictive regression using a GARCH (1,1) model where errors are assumed to follow a *t*-student distribution.

The regressions are based on monthly data that covers the sample January 2011 to March 2022.

III. Results

We start with <u>Table 1</u> which reports some basic information on the data series, the evidence of which paves the way for a formal test of how well inflation can be predicted using oil price, diesel, premix, kerosene and motor spirit.

The first takeaway from Table 1 is about the integration property and persistency of the data, which are captured using the Narayan and Popp (NP, 2010) structural break unit root test and an AR(1) model of the variable, respectively. The NP test rejects the null hypothesis of a unit root with two structural breaks, suggesting that the variables are stationary. The AR(1) model is useful to judge persistency particularly for stationary variables because stationarity is not a sufficient condition for low persistency and if persistency in predictor variables is high, predictive models need to account for it. For all the predictor variables, we find the persistency to be high, in the 0.92 to 0.98 range, with a t-statistic in the 26.73 to 48.22 range.

The endogeneity test reveals that we can only reject the null hypothesis that the slope coefficient in a regression of the residuals from a regression of *INF* on one period lagged *Price* variable on the residuals from a 6th order autoregressive model of the *Price* variable can only be rejected in the case of oil price (t-stat. = 2.75) but not for the other constituents of oil price.

The heteroskedasticity test is based on a 6th order autoregressive model of the *Price* variable and *INF*. The null hypothesis of no heteroskedasticity is based on the residuals of this model. Except for the *INF* model, the null hypothesis of no heteroskedasticity is strongly rejected for all *Price* variable-based models. This result implies that all *Price* variables are heteroskedastic.

In our first regression, we estimate if oil prices predict Fiji's inflation. The basic model results are as follows:

$$INF_t = egin{pmatrix} 0.03\ (0.14) \end{pmatrix} + egin{pmatrix} 0.001*OP_{t-1}\ (0.70) \ + egin{pmatrix} 0.01*\Delta OP_{t-1}\ (2.32)^{**} \end{pmatrix} + e_{OP,t} \end{pmatrix}$$

In this predictive regression, we find that null hypothesis of no predictability cannot be rejected as the OP variable carries a t-statistic of only 0.70. Thus, OP cannot predict Fiji's inflation. An additional possibility for predictability could be that while one-month lagged OP does not predict inflation, this evidence may change at higher lags. To explore this possibility, we estimate a 6-lag model and obtain the following result:

$$\begin{split} INF_t &= \begin{pmatrix} 0.02\\ (0.09) \end{pmatrix} + \begin{pmatrix} -0.001 * OP_{t-1}\\ (-0.26) \end{pmatrix} + \begin{pmatrix} -0.00 * OP_{t-2}\\ (-0.07) \end{pmatrix} \\ &+ \begin{pmatrix} 0.00 * OP_{t-3}\\ (0.44) \end{pmatrix} + \begin{pmatrix} -0.00 * OP_{t-4}\\ (-0.04) \end{pmatrix} \\ &+ \begin{pmatrix} 0.002 * OP_{t-5}\\ (0.18) \end{pmatrix} + \begin{pmatrix} -0.002 * OP_{t-6}\\ (-0.44) \end{pmatrix} \\ &+ \begin{pmatrix} 0.01 * \Delta OP_{t-1}\\ (2.28)^{**} \end{pmatrix} + e_{OP,t} \end{split}$$

From this predictive regression model, we see that the null hypothesis of no predictability cannot be rejected at any of the six lags, suggesting that oil prices do not have any predictive power to influence Fiji's inflation. This stands

Table 1.	Selected	summary	statistics	on	the	data
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	Oil Price	Diesel	Premix	Kerosene	Motor spirit	Inflation
NP unit root test						
Mean	149.1	1.92	2.08	1.58	2.23	0.19
SD	45.05	0.32	0.36	0.30	0.30	0.78
AR(1)	0.97 (39.44)	0.98 (48.22)	0.97 (44.12)	0.92 (26.73)	0.97 (39.02)	-0.03 (-0.34)
Endogeneity	0.014 (2.75)	-0.21 (0.23)	-0.089 (-0.12)	0.38 (0.66)	-0.30 (-0.37)	NA
ARCH(1)	182.7 (0.00)	376.7 (0.00)	277.7 (0.00)	265.0 (0.00)	253.11 (0.00)	0.22 (0.64)
ARCH(6)	40.46 (0.00)	63.07 (0.00)	46.67 (0.00)	90.30 (0.00)	40.80 (0.00)	0.86 (0.53)

This table reports selected statistics on the data that are relevant for the predictive model setup. The first test is about the integration property, which is established using the Narayan and Popp (2010) structural break unit root test and the first order autoregressive model (estimated using ordinary least squares). The t-statistic for the AR(1) coefficient, testing the null hypothesis that the coefficient is zero is reported in parenthesis. The variables mean and standard deviation (SD) are also noted. A test for endogeneity is obtained the regressing the residuals from a regression of *INF* on one period lagged *Price* variable on the residuals from a 6th order autoregressive model of the *Price* variable, where the price variables are oil price, diesel, premix, kerosene, and motor spirit. The null hypothesis of no endogeneity is evaluated based on the t-statistic reported in parenthesis. Finally ARCH(1) and ARCH(6) are autoregressive conditional heteroskedasticity tests based on a 6th order AR model of the variable.

in contrast to the literature from other countries where oil price is found to predict inflation. That oil price does not predict inflation may simply mean that there maybe other constituents of oil, such as diesel, premix, kerosene and motor spirit, that may be more relevant to inflation in Fiji. To explore this possibility, the next predict regression takes diesel prices as a predictor variable and obtains the following result:

$$egin{aligned} INF_t = egin{pmatrix} 0.15\ (0.38) \end{pmatrix} + egin{pmatrix} 0.02 * DIS_{t-1}\ (0.09) \end{pmatrix} \ + egin{pmatrix} -0.34 * \Delta DIS_{t-1}\ (-0.37)^{**} \end{pmatrix} + e_{DIS,t} \end{aligned}$$

The null hypothesis that diesel does not predict Fiji's inflation cannot be rejected with a t-statistic of 0.09. We then test if predictability has a longer horizon effect.

$$\begin{split} INF_t &= \begin{pmatrix} 0.24 \\ (0.59) \end{pmatrix} + \begin{pmatrix} -0.47*DIS_{t-1} \\ (-0.53) \end{pmatrix} + \begin{pmatrix} 2.45*DIS_{t-2} \\ (1.87) \end{pmatrix} \\ &+ \begin{pmatrix} -2.02*DIS_{t-3} \\ (-2.23) \end{pmatrix} + \begin{pmatrix} -0.18*\Delta DIS_t \\ (-0.18)^{**} \end{pmatrix} + e_{DIS,t} \end{split}$$

From this regression, we find that the null hypothesis that diesel does not predict inflation is rejected at the 6% level (with a t-stat = 1.87) at the second month, but part of this positive effect on inflation is reversed by the third month (-2.02, t-stat. = -2.23). This means that with every dollar (FJD) increase in diesel, inflation at t-2 increases by 2.45% and declines by 2.02% in month t-3. This represents a net inflationary effect of 0.43% due to every dollar increase in diesel price and this effect is felt in the first 3 months of a price increase.

Next, we consider the predictive strength of premix. Using its price as predictor, we obtain the following result:

$$INF_t = egin{pmatrix} 0.14 \ (0.37) \end{pmatrix} + egin{pmatrix} 0.02*PM_{t-1} \ (0.11) \end{pmatrix} \ + egin{pmatrix} -0.18*\Delta PM_{t-1} \ (-0.23)^{**} \end{pmatrix} + e_{PM,t} \end{cases}$$

$$egin{aligned} INF_t = egin{pmatrix} 0.05 \ (0.11) \end{pmatrix} + egin{pmatrix} -1.42*PM_{t-1} \ (-1.29) \end{pmatrix} + egin{pmatrix} 2.27*PM_{t-2} \ (1.71) \end{pmatrix} \ + egin{pmatrix} 0.08*\Delta PM_t \ (0.10) \end{pmatrix} + e_{PM,t} \end{aligned}$$

From these regressions, we see that the null hypothesis of no predictability is only rejected at the second month, suggesting that every one dollar increase in the price of premix predicts at increase in inflation by 2.27%.

The next objective is to estimate the predictive power of kerosene and the following results are found:

$$\begin{split} INF_t &= \begin{pmatrix} -0.12\\ (-0.39) \end{pmatrix} + \begin{pmatrix} 0.20*KER_{t-1}\\ (0.94) \end{pmatrix} \\ &+ \begin{pmatrix} 0.29*\Delta KEAR_{t-1}\\ (0.45) \end{pmatrix} + e_{KER,t} \\ INF_t &= \begin{pmatrix} -0.01\\ (-0.03) \end{pmatrix} + \begin{pmatrix} -0.25*KER_{t-1}\\ (-0.35) \end{pmatrix} + \begin{pmatrix} 0.75*KER_{t-2}\\ (0.80) \end{pmatrix} \\ &+ \begin{pmatrix} 0.75*KER_{t-3}\\ (0.73) \end{pmatrix} + \begin{pmatrix} -1.09*KER_{t-4}\\ (-0.94) \end{pmatrix} \\ &+ \begin{pmatrix} 0.73*KER_{t-5}\\ (0.68) \end{pmatrix} + \begin{pmatrix} -0.78*KER_{t-6}\\ (-1.26) \end{pmatrix} \\ &+ \begin{pmatrix} 0.38*\Delta KER_{t-1}\\ (0.61)^{**} \end{pmatrix} + e_{KER,t} \end{split}$$

We see that neither in the basic model nor in the augmented 6-month lag model there is any evidence that kerosene price predicts Fiji's inflation.

The final model tests the predict power of motor spirit prices in predicting Fiji's inflation rate and the following results are obtained:

$$INF_t = egin{pmatrix} 0.13 \ (0.28) \end{pmatrix} + egin{pmatrix} 0.02 * MS_{t-1} \ (0.10) \end{pmatrix} + egin{pmatrix} -0.41 * \Delta MS_{t-1} \ (-0.49) \end{pmatrix} + e_{MS,t}$$

$$\begin{split} INF_t &= \begin{pmatrix} 0.16\\ (0.29) \end{pmatrix} + \begin{pmatrix} -1.27*MS_{t-1}\\ (-1.12) \end{pmatrix} + \begin{pmatrix} 2.71*MS_{t-2}\\ (1.98) \end{pmatrix} \\ &+ \begin{pmatrix} -1.86*MS_{t-3}\\ (-1.48) \end{pmatrix} + \begin{pmatrix} 0.41*MS_{t-4}\\ (0.25) \end{pmatrix} \\ &+ \begin{pmatrix} 0.52*MS_{t-5}\\ (0.33) \end{pmatrix} + \begin{pmatrix} -0.52*MS_{t-6}\\ (-0.59) \end{pmatrix} \\ &+ \begin{pmatrix} -0.43*\Delta MS_{t-1}\\ (-0.46)^{**} \end{pmatrix} + e_{MS,t} \end{split}$$

From these results, we find predictability at the second month horizon. That is, at t-2, we find that motor spirit prices positive predict Fiji's inflation with an effect of 2.71 (t-statistic = 1.98). This implies that for every one dollar increase in motor spirit prices, inflation increases by 2.71% and this result is statistically significant at the 5% level.

IV. Concluding Remarks

This note explores the role of oil prices and its constituents (namely, diesel, premix, kerosene, and moto spirit) in predicting Fiji's inflation rate. The literature has documented the importance of oil prices to inflation. We should that while oil prices may predict inflation for larger economies, as the literature has shown, this is not true for small island states such as Fiji. For such economies, it is the other constituents of the oil market that matter. In Fiji's case, we find that oil price does not predict inflation, but diesel, premix and motor spirits predict inflation by between 0.43% to 2.71%.

Submitted: October 05, 2024 AEDT. Accepted: January 28, 2025 AEDT. Published: April 30, 2025 AEDT.



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